

About a Combinatorial Probability Limit

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Abstract. The *binomial distribution* $\mathcal{B}(n, p_n)$, for $p_n \in [0, 1]$, $(\forall) n \in \mathbb{N}$, and respectively *Poisson* (λ) *distribution*, for $\lambda \in \mathbb{R}_+^*$ are integrant parts of *Combinatorial Probabilities* (see [1]). The present note remembers us that in some given conditions, $\mathcal{B}(n, p_n)$ goes to Poisson (λ), this property being frequently applied in the quotidian life, when are used events whose appearance is rare.

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1 Introduction

Binomial distribution for a random variable X (see [3], [4]) was introduced in 1735, by the Swiss mathematician Jacques Bernoulli, and it is given by:

$$X : \left(\binom{n}{k} p^k q^{n-k} \right), \quad k \in \{0, 1, \dots, n\}, p, q \in [0, 1], p + q = 1.$$

Generally, this distribution is noted $\mathcal{B}(n, p)$ and for a some random variable X , we note $X \in \mathcal{B}(n, p)$.

Poisson (λ) *distribution* for a random variable Y (see [3], [4]), known as *the rare events law*, was introduced in 1837 by the French mathematician Siméon Poisson. So, for a same random variable $Y \in \text{Poisson}(\lambda)$, this is given by:

$$Y : \left(\frac{\lambda^k}{k!} e^{-\lambda} \right), \quad \lambda \in \mathbb{R}_+^*, k \in \mathbb{N}.$$

2 Main Result

As a preliminary fact, we compute the *average* for a random variable, *binomial distributed*.

Proposition. *Let the random variable $X \in \mathcal{B}(n, p)$, $n \in \mathbb{N}$, $p \in [0, 1]$. Then, its average is:*

$$m = M(X) = np.$$

Proof. Using the average definition for a random variable (see [2]), we have:

$$\begin{aligned}
m = M(X) &= \sum_{k=0}^n k \binom{n}{k} p^k q^{n-k} = \sum_{k=1}^n \frac{n(n-1)\dots(n-k+1)}{(k-1)!} p^k q^{n-k} = \\
&= np \sum_{l=0}^{n-1} \frac{(n-1)(n-2)\dots(n-l)}{l!} p^l q^{n-(l+1)} = \\
&= np \sum_{l=0}^{n-1} \binom{n-1}{l} p^l q^{n-1-l} = np(p+q)^{n-1} = np. \quad \square
\end{aligned}$$

In the following, we consider the sequence $(X_n)_{n \in \mathbb{N}}$, of the *random binomial variables* and we shall study its limit behaviour.

Lemma. *Let $(X_n)_{n \in \mathbb{N}^*}$, a sequence whose terms are random variables $\mathcal{B}(n, p_n)$, with $M(X_n) = \lambda \in \mathbb{R}_+^*$ and $p_n = p(n) \in [0, 1]$, $(\forall) n \in \mathbb{N}^*$. Then, we have:*

$$\lim_{n \rightarrow \infty} \mathcal{B}(n, p_n) = \text{Poisson}(\lambda).$$

Proof. From anterior Proposition, we know that $M(X_n) = n \cdot p_n$. Because $M(X_n) = \lambda$, from the hypothesis, it follows:

$$np_n = \lambda \Leftrightarrow p_n = \frac{\lambda}{n}.$$

Therefore it occurs:

$$\begin{aligned}
\lim_{n \rightarrow \infty} \mathcal{B}(n, p_n) &= \lim_{n \rightarrow \infty} \binom{n}{k} p_n^k q_n^{n-k} = \lim_{n \rightarrow \infty} \frac{n!}{k!(n-k)!} \left(\frac{\lambda}{n}\right)^k \left(1 - \frac{\lambda}{n}\right)^{n-k} = \\
&= \frac{\lambda^k}{k!} \lim_{n \rightarrow \infty} \frac{n!}{(n-k)! n^k} \left(1 - \frac{\lambda}{n}\right)^{n-k} = \\
&= \frac{\lambda^k}{k!} \lim_{n \rightarrow \infty} \frac{n-k+1}{n} \cdot \frac{n-k+2}{n} \cdot \dots \cdot \frac{n}{n} \left(1 - \frac{\lambda}{n}\right)^{n-k} = \\
&= \frac{\lambda^k}{k!} \lim_{n \rightarrow \infty} \left[\left(1 - \frac{\lambda}{n}\right)^{-\frac{n}{\lambda}} \right]^{-\frac{\lambda(n-k)}{n}} = \frac{\lambda^k}{k!} e^{-\lambda}. \quad \square
\end{aligned}$$

Remark. The anterior Lemma is a guarantee that for little probabilities $p_n = \frac{\lambda}{n}$, $n \in \mathbb{N}^*$, on limit, the distribution $\mathcal{B}(n, p_n)$ has a behaviour as Poisson (λ) , $\lambda \in \mathbb{R}_+^*$. Usually, we say that in the anterior conditions, Poisson (λ) approximates $\mathcal{B}(n, p_n)$. We remind here, that the best approximations appear when $n \geq 30$ and $np \leq 5$ (see [1]).

A chain of Lemma applications in the *Risk Theory*, can be found in [3]. \square

References

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